

# Additivity of inhibitory effects in multidrug combinations

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**From natural ecology<sup>1-4</sup> to clinical therapy<sup>5-8</sup>, cells are often exposed to mixtures of multiple drugs. Two competing null models are used to predict the combined effect of drugs: response additivity (Bliss) and dosage additivity (Loewe)<sup>9-11</sup>. Here, noting that these models diverge with increased number of drugs, we contrast their predictions with growth measurements of four phylogenetically distant microorganisms including *Escherichia coli*, *Staphylococcus aureus*, *Enterococcus faecalis* and *Saccharomyces cerevisiae*, under combinations of up to ten different drugs. In all species, as the number of drugs increases, Bliss maintains accuracy while Loewe systematically loses its predictive power. The total dosage required for growth inhibition, which Loewe predicts should be fixed, steadily increases with the number of drugs, following a square-root scaling. This scaling is explained by an approximation to Bliss where, inspired by R. A. Fisher's classical geometric model<sup>12</sup>, dosages of independent drugs add up as orthogonal vectors rather than linearly. This dose-orthogonality approximation provides results similar to Bliss, yet uses the dosage language as in Loewe and is hence easier to implement and intuit. The rejection of dosage additivity in favour of effect additivity and dosage orthogonality provides a framework for understanding how multiple drugs and stressors add up in nature and the clinic.**

In both nature and the clinic, cells are often exposed to combinations of multiple stresses and drugs. In natural ecosystems, such as the soil, dozens of microbial species capable of producing different antimicrobial compounds coexist in very close proximity, thus exposing each other to a mixture of multiple stressors<sup>1-4</sup>. In clinical settings, drug combinations, aimed at reducing side effects and counteracting resistance<sup>13-17</sup>, are becoming increasingly important in treatment for infectious diseases and cancer<sup>5-8,18,19</sup>. It is therefore of wide importance to understand how cell growth is affected by combinations of a multitude of stressors, and thereby the general rules of high-dimensionality drug arithmetic.

When combined together, drugs can interact to synergize or antagonize each other's effects relative to a null additive model. Synergy occurs when the combined effect of drugs is larger than expected based on their individual effects. Conversely, drugs can also antagonize each other, leading to a combined effect that is smaller than expected. These interactions are important in clinical settings as a way of increasing treatment potency and selectivity<sup>20-22</sup> or slowing selection for resistance<sup>14,17,23</sup>. Importantly, the definition of both synergy and antagonism relies on comparing the effect of drug combinations with a null model of 'additive expectation'<sup>6,24-27</sup>.

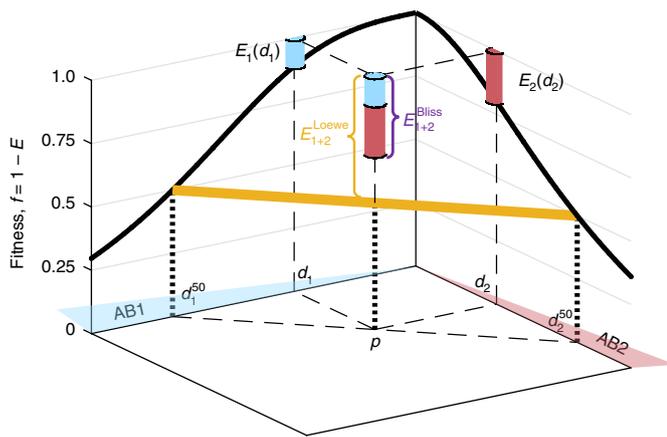
There are two primary models for the null effect of drug combinations<sup>6,28</sup>: the Bliss model<sup>10,29</sup>, which assumes response additivity; and the Loewe model<sup>9,11</sup>, which assumes dose additivity. According

to Bliss, the combined effect of two drugs  $E_{1+2}$  is simply the sum of their individual effects<sup>30</sup>,  $E_{1+2} = E_1 + E_2$ , where  $E_i = (g_0 - g_i)/g_0$  is the effect of drug  $i$  on the normalized growth rate  $g/g_0$  (Fig. 1; when effects are measured on the basis of total yield rather than growth rate, Bliss additivity becomes multiplicativity; Supplementary Note 1). In contrast, according to Loewe additivity, the effect of drugs in combination is determined not by the sum of their normalized effects, but rather by the sum of their normalized dosages, such that their combined effect is the same across all combinations that have the same total normalized dosage. Namely, according to Loewe, lines of equal combination effect in drug-dosage space (isoboles) are linear<sup>26</sup> (Fig. 1). For example, if two drugs are additive with respect to Loewe, their 50% inhibition isobole is a straight line satisfying  $d_1/d_1^{50} + d_2/d_2^{50} = 1$ , where  $d_i$  is the dose of drug  $i$  and  $d_i^{50}$  is the dose at which drug  $i$  alone causes 50% growth inhibition ( $IC_{50}$ ). Although conceptually different from one another, mechanistic support is available for both the Bliss and the Loewe models<sup>31</sup> and there is no agreement on which model should generally be used<sup>32-35</sup>. Models that implement pairwise interaction data as well as higher-order interactions can improve multidrug predictions of either Bliss or Loewe<sup>21,36-39</sup>. Yet, regardless of pairwise interactions, it remains unknown which of these two null models best predicts the combined effect of multiple drugs.

Here, measuring bacterial response to antibiotic combinations, we contrast the Bliss and Loewe models for an increasing number of drugs, where we show that expectations of these models increasingly diverge. We focus on four very different organisms: *Escherichia coli* as a model Gram-negative bacterium, *Saccharomyces cerevisiae* as a model eukaryote and the clinically important Gram-positive pathogens *Staphylococcus aureus* and *Enterococcus faecalis*. Quantifying their response to combinations of up to ten different drugs, we find that the Bliss model maintains accuracy with increased number of drugs, while the Loewe model loses its predictive power. Indeed, in contrast to Loewe, which predicts that the total drug dosage required for inhibition is constant, we find that total dosage increases monotonically with the number of drugs. Interestingly, our data show that this increase follows a square-root scaling, inspiring a simple model for orthogonality of dose additivity that follows a classical evolutionary optimization principle developed by R. A. Fisher<sup>12</sup>.

To contrast the Bliss and Loewe models, we calculated how their predictions scale with increased number of drugs. As a natural measure of the combined potency of multiple drugs, we considered the total drug dosage needed to achieve a given level of inhibition. Defining 'total dosage' ( $D$ ) as the sum of the concentration of the individual drugs weighted by their corresponding  $IC_{50}$  values,  $D = \sum_i d_i / d_i^{50}$ , the 'combination potency' ( $D^{50}$ ) is the total dosage  $D$  that yields 50% growth inhibition. As the number of drugs  $N$  increases, Loewe additivity predicts that the combination

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**Fig. 1 | Schematic depiction of effect additivity (Bliss) and dosage additivity (Loewe).** Given fitness as a function of dosage of each of the individual drugs ( $f_i = 1 - E_i$ , dose response curve, black solid line), Bliss and Loewe models predict the fitness  $f_{1+2} = 1 - E_{1+2}$  at any given point  $P = [d_1, d_2]$  in the drug concentration space. The Bliss prediction assumes additivity of normalized drug effects,  $E_{1+2}^{\text{Bliss}} = E_1 + E_2$ , where  $E_1$  and  $E_2$  are the individual drug effects at their cognate concentrations (cyan and red piles, respectively). The Loewe model, on the other hand, assumes additivity of normalized drug dosage, such that the combined drug effect is fixed along linear lines of constant total normalized dosage (yellow,  $E_{1+2}^{\text{Loewe}}$  equals 50% in the example point  $P$ , and more generally is given by solving for  $E$ :  $d_1/d_1^E + d_2/d_2^E = 1$ , where  $d_i^E$  is the concentration of drug  $i$  that leads to inhibition level  $E$ ).

potency remains fixed,  $D_{\text{Loewe}}^{50}(N) = 1$ . The prediction of Bliss, on the other hand, depends on the individual dose response curves of the different drugs. Assuming a Hill equation<sup>40</sup> for the single-drug dose response  $E_i(d_i) = 1/[1 + (d_i^{50}/d_i)^h]$ , where  $h$  is the Hill coefficient, and equating the Bliss prediction of the combined effect  $E_{1..N} = \sum_{i=1..N} E_i$  to 50%, yields the Bliss-predicted scaling of combination potency with the number of drugs:  $D_{\text{Bliss}}^{50}(N) = N(2N - 1)^{-(1/h)}$ . Thus, while Loewe predicts that the total dosage required for inhibition is constant, Bliss predicts that it increases with the number of drugs (assuming Hill coefficient,  $h > 1$ ). The two models can therefore best be contrasted by measuring the combined action of increased number of drugs.

Starting with *E. coli*, we considered ten mechanistically different antibiotics and measured their growth-inhibitory effect individually as well as in combinations with increased number of drugs. We chose antibiotics acting on a range of cellular functions, including cell wall synthesis, DNA replication, transcription and translation (Supplementary Table 1). Measuring optical density (OD) versus time of bacterial growth on gradients of each of the individual drugs, we determined the dose response curve  $g(d_i)$  for each of the drugs (Fig. 2a and Supplementary Fig. 1). These dose response curves are well fitted by Hill functions, defining the concentrations  $d_i^{50}$  of 50% inhibition for each of the drugs in isolation (Supplementary Fig. 13a,b).

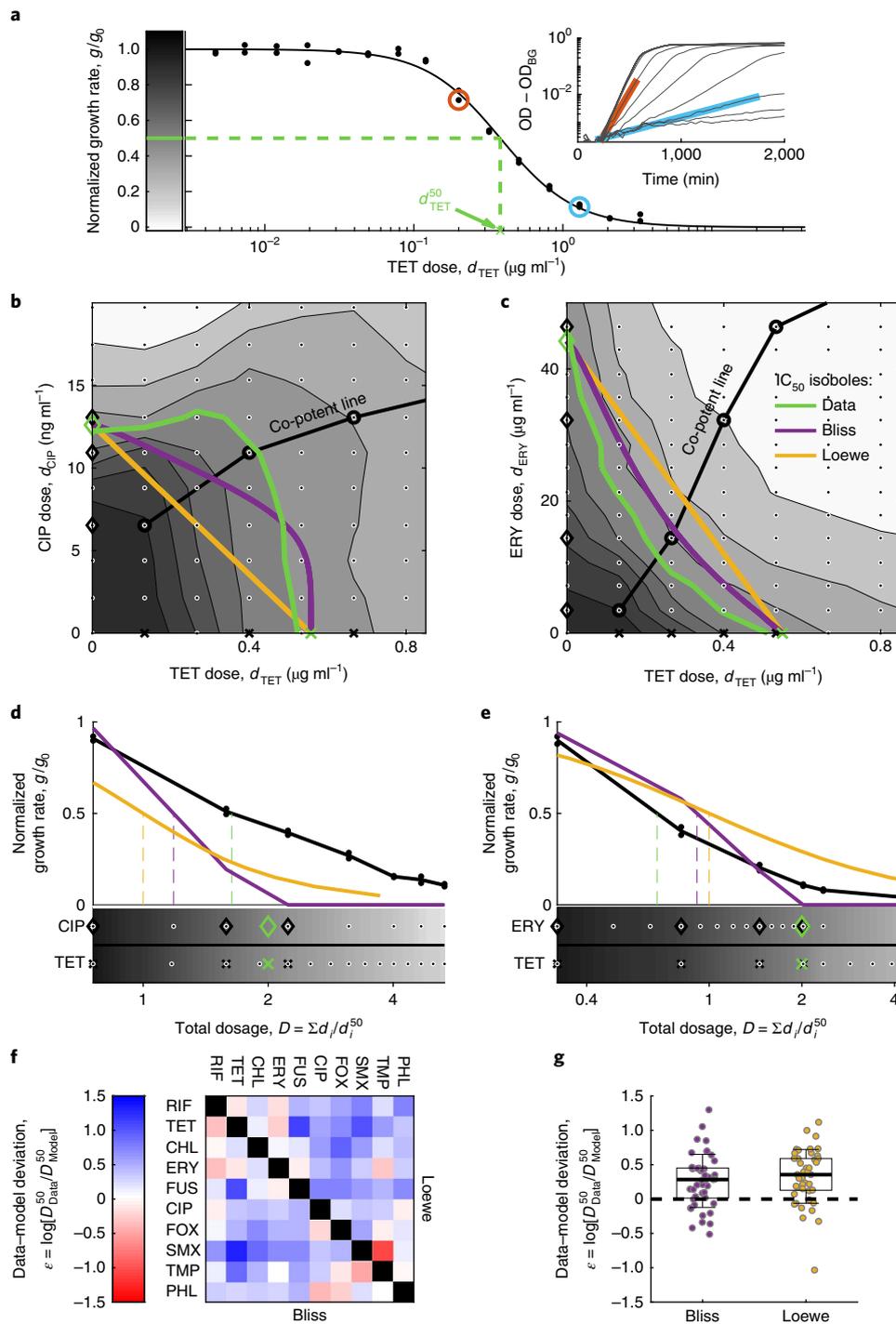
Moving to drug pairs, we measured their combination potency and compared it to Bliss and Loewe predictions. We first measured the full response surface across two-dimensional dose gradients for two drug pairs: tetracycline and ciprofloxacin (TET-CIP) and tetracycline and erythromycin (TET-ERY), representing well-known examples of antibiotic antagonism and synergy (Fig. 2b,c response surface and  $IC_{50}$  isoboles; Supplementary Fig. 12a)<sup>41,42</sup>. Using the growth measurements of the individual drug gradients  $E_i(d_i)$ , we derived the response surface and the  $IC_{50}$  isobole predictions of Loewe (straight line connecting the points  $[d_1^{50}, 0]$  and  $[0, d_2^{50}]$ )

and Bliss (the set of all points  $[d_1, d_2]$  satisfying  $E_1(d_1) + E_2(d_2) = 50\%$ , Methods). As expected, the measured  $IC_{50}$  isobole lies above these predictions for the TET-CIP pair (indicating antagonism) and below for the TET-ERY pair (indicating synergy). While these two-dimensional gradients allow clear definition of synergy and antagonism, they require many growth measurements and become combinatorially prohibitive in a high-dimensional multidrug space.

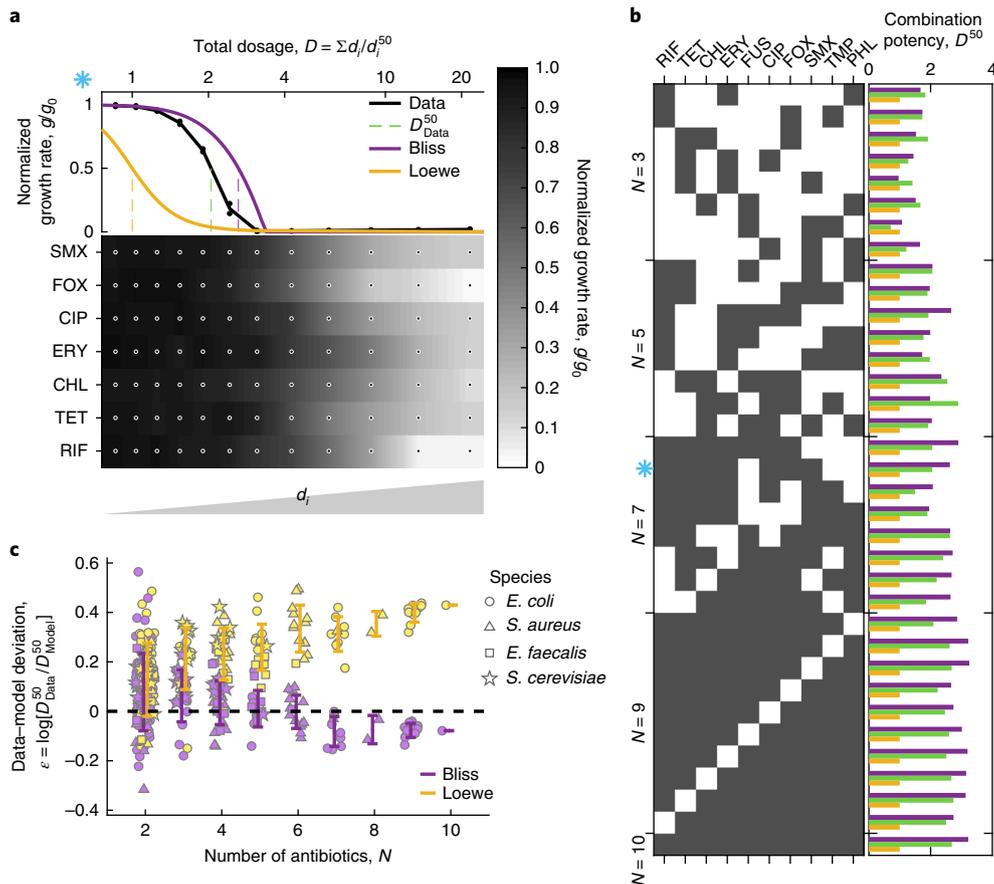
To effectively sample the concentration space of multiple drugs, we performed growth measurements along a ‘co-potent’ line<sup>38</sup>, a curve in concentration space where the individual drugs have equal potencies in isolation ( $E_1(d_1) = E_2(d_2) = \dots = E_N(d_N)$ , Fig. 2b–e). This co-potent line sampling method vastly reduces the dimensionality of the required measurements while guaranteeing that null models are evaluated in a region in drug concentration space where all drugs are active, rather than one in which the combined effect is dominated by a subset of drugs. Identifying the point  $P = (d_1, d_2, \dots, d_N)$  on the co-potent line where growth is inhibited by 50% yields the combination potency,  $D_{\text{Data}}^{50} = \sum_i d_i/d_i^{50}$ . This measured combination potency was contrasted with the expected potencies  $D_{\text{Bliss}}^{50}$  and  $D_{\text{Loewe}}^{50}$ , defined as the points along the co-potent line where the single-drug-based calculations of Bliss and Loewe predict 50% inhibition (Methods). The interaction between drugs was then defined as the deviation between the observed and predicted potencies of the combination  $\varepsilon = \log(D_{\text{Data}}^{50}/D_{\text{Model}}^{50})$ , which captures the extent of antagonism ( $\varepsilon > 0$ ) and synergy ( $\varepsilon < 0$ ) (Fig. 2d,e).

Measuring combination effects for all drug pairs, we find that their joint potencies are similarly well predicted by both the Bliss and the Loewe models. For each of the 45 drug pairs, we measured their dose response along co-potent concentration gradient and determined their combination potency,  $D_{\text{Data}}^{50}$  (Methods, Fig. 2f and Supplementary Figs. 12b and 13a). Comparing these combination potencies with predictions of the Bliss and Loewe null models, we find that both positive ( $\varepsilon > 0$ , antagonism) and negative ( $\varepsilon < 0$ , synergism) deviations are prevalent with respect to either model (Fig. 2f,g). This prevalence of both antagonism and synergy among drug pairs overwhelms any small deviations between the two models (Fig. 2g;  $\sigma(\varepsilon_{\text{Bliss}}) = 0.39$ ;  $\sigma(\varepsilon_{\text{Loewe}}) = 0.39$ ;  $\langle \varepsilon_{\text{Bliss}} \rangle = -\langle \varepsilon_{\text{Loewe}} \rangle = -0.066$ ,  $t$ -test:  $P = 0.42$ ). Further, clustering drugs on the basis of these pairwise interactions, defined with respect to either Bliss or Loewe, leads to similar grouping by mechanism of action (Supplementary Fig. 2; possible small advantage to Bliss in resolving fine functionality differences)<sup>43</sup>. The similarity of pairwise null predictions, the prevalence and magnitude of pairwise interactions with respect to both models, and their similar correlation with cellular function prohibit distinction of the Loewe and Bliss null models based on drug pairs.

However, for increasing number of drugs, we find that the combined effect is well predicted by Bliss, while the Loewe prediction systematically diverges. Given that predictions of the two models should diverge with increased number of drugs, we measured the combined effect of multiple combinations with a varying number of antibiotics. We chose 35 combinations of 3 to 10 antibiotics, including 8 randomly chosen sets from each combination size of  $N = 3, 5$  and 7 drugs, all 10 sets of 9 drugs, and the whole 10-drug set (Fig. 3a,b; Supplementary Figs. 3a, 12c and 13b and Supplementary Tables 9 and 10). Following the procedure used for the drug pairs, we measured the combined effect of each multidrug set as a function of total dosage along co-potent lines and identified their combination potencies  $D^{50}$ . Contrasting these measured potencies with the predicted potencies of Bliss and Loewe based on the single-drug measurements, we find that the Bliss model maintains good accuracy regardless of the number of drugs, while the accuracy of the Loewe model declines as the number of drugs increases (Fig. 3c, dots, and Supplementary Figs. 4 and 5). This rejection of Loewe in favour of Bliss is independent of the co-potency of the drug mixture (Supplementary Fig. 6) and also appears when using the



**Fig. 2 | Pairwise measurements do not resolve the Bliss and Loewe models of additivity.** **a**, Representative single-drug dose response curve showing the normalized growth rate  $g/g_0$  along a concentration gradient of TET (black dots, replicates), the Hill equation fit ( $g/g_0 = 1/(1 + (d_{TET}/d_{TET}^{50})^{h_{TET}})$ , black line) and the  $IC_{50}$  ( $d_{TET}^{50}$ , green dashed line). Inset: growth rates  $g$  were calculated by fitting  $OD_{600nm}$  measurements over time (black) to exponential function,  $OD = OD_0 \cdot 2^{g \cdot t} + OD_{BG}$  (cyan and red). **b, c**, Response surface showing growth rates (greyscale indicated in **a**) over a two-dimensional grid gradient (dots) of the antagonistic antibiotic pair TET-CIP (**b**) and the synergistic pair TET-ERY (**c**). The measured  $IC_{50}$  isoboles (green) are contrasted with Bliss (purple) and Loewe (orange) predictions. Indicated are the co-potent lines (circles), the corresponding co-potent single drugs (black crosses for TET, black diamonds for CIP and ERY) and the  $IC_{50}$  values (green symbols). **d, e**, Dose response along the co-potent line of the two drug mixtures (TET-CIP, **d**; TET-ERY, **e**) as a function of total dosage  $D = d_1/d_1^{50} + d_2/d_2^{50}$ . Measured normalized growth rates of the combined drugs are contrasted with Bliss and Loewe predictions based on the single-drug data (shown below). All symbols correspond to those in **b** and **c**. **f, g**, Data-model deviations ( $\epsilon = \log(D_{Data}^{50}/D_{Model}^{50})$ ), indicates the difference between measured ( $D_{Data}^{50}$ ) and predicted ( $D_{Bliss/Loewe}^{50}$ ) combination potencies) for each of the two models are presented as an interaction matrix (**f**) and a box plot (**g**, box at first and third quartiles, whiskers at mean  $\pm$  two standard deviations, black dashed line represents perfectly accurate prediction). There is no significant difference between the models in their predictions of measured pairwise potencies (45 different combinations, two-sided  $t$ -test,  $P = 0.42$ ).



**Fig. 3 | The Loewe model of additivity loses its predictive power with increased number of combined drugs.** **a**, An example of dose response along the co-potent line of a mixture of seven drugs. The measured normalized growth rate calculated on the basis of duplicate measurements (dots) and the combined potency  $D_{Data}^{50}$  of the drug mixture are contrasted with Bliss and Loewe predictions calculated on the basis of the single-drug measurements (below, greyscale). **b**, The combination potency ( $D_{Data}^{50}$ , green) is contrasted with predictions of Bliss ( $D_{Bliss}^{50}$ , purple) and Loewe ( $D_{Loewe}^{50}$ , yellow) for 35 different combinations of 3 to 10 drugs (black squares). The blue asterisk matches the combination shown in **a**. **c**, The deviation of each of the models from the data  $\varepsilon = \log[D_{Data}^{50}/D_{Model}^{50}]$  is plotted for 172 different combinations as a function of the number of drugs (error bars represent one standard deviation of each combination size), showing that the Loewe predictions deviate from the data with increased number of drugs, while Bliss predictions remain accurate (the dashed line represents perfectly accurate prediction).

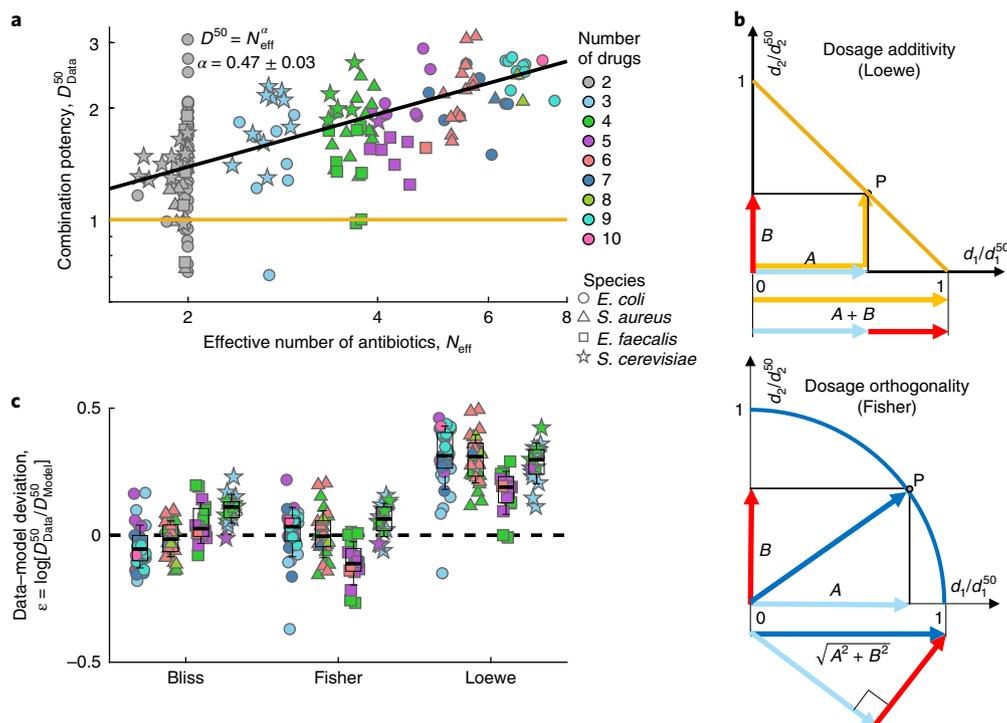
multiplicative form of the Bliss model (more suitable for yield measurements, Supplementary Note 1; Supplementary Fig. 7), as well as when considering alternative derivation of growth rates from OD curves (Supplementary Fig. 1c)<sup>34</sup>. Our *E. coli* growth data therefore show that the effect of combinations of many diverse drugs is well predicted by Bliss, and not by Loewe.

To test the generality of these findings, we applied our methodology to the Gram-positive pathogens *E. faecalis* and *S. aureus* as well as to the eukaryotic model of *S. cerevisiae*. For each of these organisms, we chose a repertoire of diverse drugs and measured their combined effect in random sets of increased number of drugs (up to eight in *S. aureus*, six in *E. faecalis* and five in *S. cerevisiae*; Supplementary Tables 1, 9 and 10 and Supplementary Figs. 3b–e, 12d–g and 13c–f). As in *E. coli*, both the Bliss and Loewe models provided comparable predictions for combinations of small number of drugs and when the number of drugs increases the fitness was well predicted by Bliss while Loewe's prediction diverged (Fig. 3c and Supplementary Figs. 4 and 5). Furthermore, the power of Bliss over Loewe is also seen for combinations involving strongly synergistic drug pairs, treating the pair mixture as a single agent (TMP–SLF in *S. aureus*; Supplementary Fig. 8). We conclude that across species the Loewe model, predicting that the total dosage required for inhibition is fixed regardless of the number of drugs,

can be rejected as a general predictor for the potency of diverse multidrug combinations.

Next, we tested how the potency of drug combinations, namely the total dosage required for inhibition, varies with the number of drugs. To account for any slight experimental deviations from the ideal co-potent line, we use a natural entropy-like definition of an effective number of drugs  $N_{\text{eff}}$  that is based on the uniformity of the individual drug effects ( $N_{\text{eff}}$  equals  $N$  if all drugs have the same effect; is slightly smaller than  $N$  when these effects are uneven; and converges to 1 at the extreme case when a single drug dominates; see definition of  $N_{\text{eff}}$  in Fig. 4 caption). Contrary to the Loewe prediction, we find that the total dosage required for inhibition increases with the effective number of drugs (Fig. 4a and Supplementary Fig. 8a). Moreover, this inhibitory total dosage seems to obey a simple scaling law: it increases as the square root of the effective number of drugs ( $D^{50} = (N_{\text{eff}})^{\alpha}$ , least-square fit yields:  $\alpha = 0.47 \pm 0.03$ ).

The square-root scaling of the inhibitory dosage with number of drugs can be explained by an approximation to Bliss, inspired by the classical optimization principle of Fisher's geometric model of adaptation<sup>12</sup>. Fisher's model describes the fitness  $f$  in a space of  $N$  independent orthogonal phenotypes and assumes that it declines as a function of the Euclidean distance from an optimal point ( $f = e^{-r^2}$ , where  $r^2 = \sum_{i=1..N} x_i^2$  and  $x_i$  are phenotypic distances from the



**Fig. 4 | A square-root scaling law of inhibitory total dosage with effective number of drugs is explained by a simple dosage-orthogonality model. a**, The combination potency,  $D_{Data}^{50}$  of all 172 different drug combinations is plotted as a function of effective number of drugs ( $N_{eff} = \exp[-\sum_i p_i \log(p_i)]$ , where  $p_i = E_i / \sum_j E_j$ , and  $E_i(d_i)$  are the single-drug individual effects at their cognate concentrations; colours represent the actual number of drugs,  $N$ ). In contrast to Loewe, which assumes that the total dosage required for inhibition is fixed (yellow line), the total dosage increases as the square root of the effective number of drugs (black line, fit of  $D^{50} = (N_{eff})^\alpha$  yields  $\alpha = 0.47 \pm 0.03$ , 0.95 confidence interval). **b**, The square-root scaling is explained by a Fisher-inspired dose-orthogonality model, which assumes that for small perturbations the dosages of independent drugs should be added as orthogonal vectors rather than linearly as in Loewe. Hence, isoboles of  $X\%$  inhibition are spherical surfaces defined by  $\sum_i (d_i/d_i^{50})^2 = 1$  (Fisher, bottom, circles in two-drug space, blue line) instead of linear surfaces (Loewe, top, straight lines in two-drug space, yellow line). **c**, Data-model deviation  $\epsilon = \log(D_{Data}^{50}/D_{Model}^{50})$  of all the combinations of more than two drugs for each of the three models (box at median and first and third quartiles, whiskers at mean  $\pm$  two standard deviations, black dashed line represents perfectly accurate prediction). The data strongly reject Loewe and are instead consistent with both Fisher and Bliss (96 data points, two sided  $t$ -test: Loewe,  $P < 10^{-40}$ ; Bliss,  $P = 0.61$ ; Fisher,  $P = 0.8$ ).

optimal point). For a given fitness value, the phenotypic distances  $x_i$ , therefore decline as  $1/\sqrt{N}$  and their sum,  $D = \sum_{i=1}^N x_i^2$ , increases as  $\sqrt{N}$ . The analogy of drug dosages with Fisher's phenotypes, and more generally the analogy of drug and mutations, explains the square-root scaling of total inhibitory dosage with number of drugs and underscores that drug concentrations should be summed not linearly by simple addition as in Loewe, but rather as the geometric sum of orthogonal vectors (Fig. 4b; of course, orthogonality is an idealization from which drug combinations can deviate due to synergy or antagonism, or when similar drugs act along the same axis). This Fisher-inspired 'dose-orthogonality' model can also be derived as a second-order approximation of Bliss additivity at the limit of small dosages, emphasizing the dependency of this approximation on optimality of the growth rate at the origin (Supplementary Note 2; for Hill coefficients close to 1, as common in the drug combinations applied to *E. faecalis*, Bliss prediction can be closer to Loewe than to Fisher). Indeed, we find that even strongly interacting drug pairs assume more circular isoboles for small fitness effects (Fig. 2b,c and Supplementary Fig. 10) and that the square-root scaling becomes more accurate at a lower level of inhibition (Supplementary Fig. 9b). Similarly to Bliss and in contrast to Loewe, combination potency predictions of this dosage-orthogonality model (derived by intersecting the co-potent lines with spherical  $IC_{50}$  isoboles  $\sum_i (d_i/d_i^{50})^2 = 1$ , Methods) are consistent with the drug combination measurements (Fig. 4c). Yet, unlike Bliss these predictions do not require fine

measurements of the minute individual drug effects  $E_i(d_i)$ , but rather depend on the more robust measurements of their individual  $IC_{50}$  dosages,  $d_i^{50}$ . While it allows the use of dose language like Loewe, the dose-orthogonality model provides an intuitive and robust approximation of Bliss's response additivity (Supplementary Table 2), which well predicts the potency of drug combinations and explains the square-root scaling law of potency with number of drugs.

Across diverse taxons, our measurements reject the Loewe model of dosage additivity for predicting the effect of combination of multiple diverse drugs, favouring the Bliss effect additivity and motivating a simple model of dosage-orthogonality. In contrast to Loewe additivity, which predicts that the total dosage required for inhibition is fixed, we find that the total inhibitory dosage increases with the number of drugs, following a square-root scaling law. This general reduction in potency with increased number of drugs could be important for understanding ecological environments where bacteria are exposed to a multitude of drugs and stresses and any one toxin does not typically work on its own but rather combined within a complex soup of natural inhibitors. While the results could also be important in the clinical settings where multiple drugs are combined, any such implications will require extending the conceptual approach and methodology and to consider the killing regime, the impact on multispecies communities and the complexity that stems from transient effects introduced by the pharmacokinetics and pharmacodynamics of each of the individual drugs. The square-root

scaling supports a model for drug additivity where dosages of independent drugs add up orthogonally rather than linearly. This dosage-orthogonality model provides an approximation to Bliss, yet it uses dosage arithmetic that allows a more robust implementation and simple intuition. It will be interesting to explore the generality of these results and the limit on the number of orthogonal pharmacological axes as more antibiotics and stresses are added, as well as beyond the minimal inhibitory concentration and in more complex systems such as in cancer therapy. Throughout such clinical systems and natural ecologies, our findings provide a uniform framework for understanding the null arithmetic of many-drug combinations.

## Methods

**Strains and media.** Experiments were performed with: *E. coli* strain MC4100 in M9 media ( $\text{Na}_2\text{HPO}_4$  6 g l<sup>-1</sup>,  $\text{Na}_2\text{HPO}_4$  3 g l<sup>-1</sup>, NaCl 0.5 g l<sup>-1</sup>,  $\text{NH}_4\text{Cl}$  1 g l<sup>-1</sup>, glucose 2 g l<sup>-1</sup>, Casamino acids 1 g l<sup>-1</sup>, thiamine 0.34 g l<sup>-1</sup>,  $\text{MgSO}_4$  1 mM,  $\text{CaCl}_2$  0.1 mM); *S. aureus* sp. Rosenbach (ATCC 29213 - Wichita) in Mueller-Hinton broth; *E. faecalis* ATCC 49757 in brain heart infusion broth; and *S. cerevisiae* BY4741 Euroscarf in YPD broth (yeast extract 1%, peptone 2%, glucose 2%). Antibiotics were added as indicated (Sigma Aldrich).

**Growth rate assay.** The data represent seven different experimental set-ups: three for *E. coli*: two-dimensional gradient of ERY-TET and CIP-TET (Fig. 2b–e and Supplementary Figs. 9 and 12a), all 45 pairwise combinations (Fig. 2f,g and Supplementary Figs. 12b and 13a), and 35 combinations of order higher than two (Figs. 3 and 4, dots, and Supplementary Figs. 12c and 13b); two for *S. aureus*: 22 combinations not involving beta-lactam antibiotics, and 21 combinations involving beta-lactam antibiotics (Figs. 3c and 4, triangles, and Supplementary Figs. 12d,e and 13c,d); one for *E. faecalis*, composed of 23 combinations (Figs. 3c and 4, squares, and Supplementary Figs. 12f and 13e); and one for *S. cerevisiae*, composed of 26 combinations (Figs. 3c and 4, stars; Supplementary Figs. 12g and 13e). In each of these experiments, an inoculum of  $10^4$  cells was inoculated into 150  $\mu\text{l}$  of media in a Nunc 96-well flat-bottom microplate. Antibiotics were added into the wells as indicated using a D300e digital dispenser (Tecan), which dispenses a nanolitre-scale volume of each antibiotic. Each concentration combination was performed in duplicate wells. Multiple untreated control wells (no antibiotics) were designated on each plate (2–6% of the wells in each experiment). To avoid systematic positional effects across the plates, the wells chosen for each drug combination on the plates were randomized. The plates were incubated at 30 °C with shaking (Liconic orbital shaker STX44), and  $\text{OD}_{600\text{nm}}$  was measured at least every 25 min using a Tecan robotic system and the Infinite M200 Pro reader. To enhance uniformity, the plate orientations in the shaker were rotated 180° following every measurement.

**Growth rate analysis.** Data analysis was performed using MATLAB. The  $\text{OD}_{600\text{nm}}$  measurements were averaged using a running window of two data points. The log phase of each of these curves was determined using an algorithm based on the sensitivity analysis method ('tornado diagram', Supplementary Fig. 1), and OD data in this constant exponential growth region were fitted to  $\text{OD} = \text{OD}_0 \cdot 2^{g \cdot t} + \text{OD}_{\text{BG}}$  (in some of the wells, the log-phase period was mis-determined, and were cured manually; all fits are indicated in Supplementary Fig. 12). From these fitted parameters, we obtained for each well the background  $\text{OD}_{\text{BG}}$  and the growth rate  $g$ . Fitness for each drug mixture was then calculated as normalized growth rates  $g/g_0$ , where  $g$  is the average growth rate in duplicate wells and  $g_0$  is the median growth rate of all untreated wells in the experiment. All calculated normalized growth rates are available in Supplementary Table 10. While OD is commonly used as a measure of microbial density, the relation between OD and cell number depends on cell size and morphology, which can be affected by the antibiotics. As the growth rate is defined as the logarithmic derivative of the OD, it is independent of this effect of antibiotic on cell size, as long as these effects are constant throughout the fitted region. Therefore, we measure the growth rate in the regime of steady exponential growth long after the addition of the antibiotic.

**Determining drug concentrations for co-potent combinations.** To achieve co-potent concentrations for the different drugs, we first measured growth rates on gradients of each of the individual drugs and identified the  $\text{IC}_{20}$  and  $\text{IC}_{80}$  for each of the drugs,  $d_i^{20}$  and  $d_i^{80}$ . Then, we designed the drug gradients such that the individual drugs have matching effects:  $d_i(k) = d_i(1) / w_i^k$ , where  $d_i(k)$  is the dosage of drug  $i$  at dilution number  $k$ ,  $w_i = \sqrt[3]{d_i^{80}/d_i^{20}}$  is the dilution factor (chosen to have three dilutions between the  $\text{IC}_{20}$  and  $\text{IC}_{80}$ ),  $d_i(1) = d_i^{80} \times w_i^3$  is the highest concentration (strong enough to inhibit growth) and  $k$  varies from 1 to 15 such that the last concentration  $d_i(15) = d_i^{80} \times w_i^{-11}$  is low enough to have negligible effect (all drug concentrations are provided in Supplementary Tables 3–8). For example: 7.3 ng ml<sup>-1</sup> and 53.1 ng ml<sup>-1</sup> of TET inhibit the growth of *S. aureus* by 20% and 80%, respectively; hence,  $w_{\text{TET}} = (53/7.3)^{1/3} = 1.93$ . At the same time, 35 ng ml<sup>-1</sup> and 152 ng ml<sup>-1</sup> ERY

inhibit the growth by 20% and 80%, respectively; hence,  $w_{\text{ERY}} = (152/35)^{1/3} = 1.64$ . On the co-potent line, we mix ERY and TET (Supplementary Fig. 3b) with the following concentrations:  $d_{\text{TET}}^{80} \cdot w_{\text{TET}}^3 = 53 \times 1.93^3 = 386 \text{ ng ml}^{-1}$  TET with  $d_{\text{ERY}}^{80} \cdot w_{\text{ERY}}^3 = 152 \times 1.64^3 = 664 \text{ ng ml}^{-1}$  Ery ('O' in Supplementary Table 5), 199 ng ml<sup>-1</sup> TET with 405 ng ml<sup>-1</sup> ERY ('N'), ..., 0.52 ng ml<sup>-1</sup> TET with 2.93 ng ml<sup>-1</sup> ERY ('E').

**Calculating model predictions for Bliss, Loewe and Fisher.** The models predict fitness across drug concentration space  $f(d_1, d_2, \dots, d_n)$  based on the single-drug dose response curves ( $f_i(d_i)$ , Hill fitted, Supplementary Fig. 12). For any given level of inhibition  $E$ , we define in each of the models the  $N-1$ -dimensional isobolic surface of all points  $(d_1, d_2, \dots, d_n)$  in concentration space predicted to have fitness  $f=1-E$ . For Bliss, this isobolic surface is the collection of points satisfying  $\sum_i (1-f_i(d_i)) = 1-E$ . For Loewe, the isobolic surface is predicted to be linear, satisfying  $\sum_i (d_i/d_i^E) = 1$ , where  $d_i^E$  is the concentration of drug  $i$  that inhibits growth by  $E$ , as defined by  $f_i(d_i^E) = 1-E$ . For Fisher, this isobolic surface is an  $N-1$ -dimensional sphere defined by  $\sum_i (d_i/d_i^E)^2 = 1$ . Substituting  $E=50\%$ , these calculations define the  $\text{IC}_{50}$  isoboles (Fig. 2b,c). The intersections of the co-potent lines with the predicted isobolic surfaces for ranges of values of  $E$  define the predicted fitness along the co-potent line (Figs. 2 and 3, Supplementary Fig. 3 and Supplementary Table 11). In particular, the intersections of the co-potent lines with the  $E=50\%$  isobolic surfaces define the predicted potencies  $D_{\text{Bliss}}^{50}$ ,  $D_{\text{Loewe}}^{50}$  and  $D_{\text{Fisher}}^{50}$  (Figs. 2–4, Supplementary Fig. 3 and Supplementary Table 11).

**Reporting Summary.** Further information on research design is available in the Nature Research Reporting Summary linked to this article.

**Code availability.** MATLAB codes to calculate growth rate and model predictions are available on the laboratory website at <https://kishony.net.technion.ac.il/resources/>; any additional codes are available from the corresponding author.

## Data availability

The data sets generated during and/or analysed during the current study are available from the corresponding author on reasonable request.

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### Competing interests

The authors declare no competing interests.

### Additional information

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No software were used.

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The data in this study were analyzed using custom MATLAB scripts. Our analysis was compared to growth rate calculated by the R-package "grofit" (version 1.1.1).

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Data exclusions	No data exclusion.
Replication	All growth rates were calculated based on duplicates. the to replicates set in different plates and positions of wells were randomized in each plate (using the d300e digital dispenser software).
Randomization	Choice of antibiotics in each group were done randomly (rand command in MATLAB limited so no drug present in 3 combinations more than the least represented) .
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